

Kalman Filters

Wednesday, April 19, 2017 9:58 AM

minimize $E\{e^2[n]\}$ ← like LMS except.

1. KF does not observe $d[n]$ directly
2. KF works in highly non-stationary environments

the KF is state-based

We have some system model and keep track of the system state

example: we have a car moving at a constant speed

state: car's position & speed

$$\underline{x}(n) = \begin{pmatrix} p(n) \\ s(n) \end{pmatrix} \begin{array}{l} \leftarrow \text{position} \\ \leftarrow \text{speed} \end{array}$$

at time $n+1$

$$\underline{x}(n+1) = \begin{pmatrix} p(n) + \Delta t \cdot s(n) \\ s(n) \end{pmatrix}$$

we can write this as

$$\underline{x}(n+1) = A \underline{x}(n) + \underline{w}(n) \quad \text{where } A = \begin{pmatrix} 1 & \Delta t \\ 0 & 1 \end{pmatrix}$$

we usually have noise

- this noise is the driver - he can adjust the speed

$$\underline{w}(n) = \begin{pmatrix} 0 \\ \text{something} \end{pmatrix}$$

$$y(n) = x_1(n) + v(n)$$

measurement noise

$$y(n) = C \underline{x}(n) + v(n)$$

$$C = \begin{pmatrix} 1 & 0 \end{pmatrix}$$

we measure position but not speed and we have measurement noise

Approximation Error

Wednesday, April 19, 2017 10:11 AM

$$\hat{X}_{\text{trunc}} - \hat{X}_{\text{pinv}} = \underbrace{\sum_{k=R'+1}^R -\langle x, v_k \rangle v_k}_{\text{approximation error}} + \underbrace{\sum_{r=1}^{R'} \frac{1}{\sigma_r} \langle e, u_r \rangle v_r}_{\text{noise error}}$$

These are mutually orthogonal

→ this means that

$$\begin{aligned} \|\hat{X}_{\text{trunc}} - \hat{X}_{\text{pinv}}\|_2^2 &= \|\text{Approx. error}\|_2^2 + \|\text{noise error}\|_2^2 \\ &= \sum_{k=R'+1}^R |\langle x, v_k \rangle|^2 + \sum_{r=1}^{R'} \frac{1}{\sigma_r^2} |\langle e, u_r \rangle|^2 \end{aligned}$$

Assuming, as before, that $e[m] \sim N(0, V^2)$

$$E[\|\text{Noise error}\|_2^2] = V^2 \sum_{r=1}^{R'} \frac{1}{\sigma_r^2}$$

Aside

$$\|x\|_2^2 = \langle x, x \rangle$$

$$\sum_{r=1}^R \frac{1}{\sigma_r} \langle e, u_r \rangle v_r$$

the error due to noise $\hat{X}_{\text{trunc}} - \hat{X}_{\text{pinv}}$

scalar $\delta_r = \frac{1}{\sigma_r} \langle e, u_r \rangle$

$$\sum_r \delta_r v_r \Rightarrow \|\text{error}\|_2^2 = \langle \sum_r \delta_r v_r, \sum_k \delta_k v_k \rangle$$

$$= \sum_r \delta_r \langle v_r, \sum_k \delta_k v_k \rangle$$

$$\sum_r \delta_r \sum_k \delta_k \langle v_r, v_k \rangle = \begin{cases} 1 & r=k \\ 0 & \text{else} \end{cases}$$

$\frac{\|u_r\|_2}{\sigma_r}$ (normalization factor)

$$\begin{aligned}
 &= \sum_r \delta_r^2 = \sum_r \frac{1}{\sigma_r^2} \langle e, u_r \rangle^2 \\
 &= \frac{1}{\sigma_R^2} \langle u_R, u_R \rangle^2 \quad \left. \begin{array}{l} \text{else} \\ \text{if } e = u_R \end{array} \right\} \text{to}
 \end{aligned}$$

Kalman Filter Notation

Wednesday, April 19, 2017 10:11 AM

State equation

$$\underline{x}(n+1) = A(n+1, n) \underline{x}(n) + B(n) \underline{w}(n)$$

Observation equation

$$y(n) = C(n) \underline{x}(n) + D(n) v(n)$$

goal is to find $x(n)$

$v(n)$ - observation noise

(for this development, assume B & D are identity matrices)

$$E\{v(n)v^T(n)\} = Q_v \delta(n-k)$$

each time step is uncorrelated with all others but at a particular time, there may be internal correlations (in Q_v)

$w(n)$ - process noise

$$E\{w(n)w^T(n)\} = Q_w \delta(n-k)$$

Let $\hat{x}(n|i)$ be the unbiased minimum mean-square estimate of $x(n)$ at time n given all the measurements up to and including time i

Let $e(n|i)$ be the corresponding state estimation error

$$e(n|i) = \underline{x}(n) - \hat{x}(n|i)$$

and

$$P(n|i) = E\{e(n|i)e^T(n|i)\}$$
 is the error covariance matrix

Goal: minimize MSE

$$E\{\|e(n|n)\|^2\} = \text{Tr}\{P(n|n)\} = \sum_{k=0}^{n-1} E\{e_k^2(n|n)\}$$

solution

$$\hat{x}(n|n) = A(n, n-1) \hat{x}(n-1|n-1) + K(n) \left[y(n) - C(n) A(n, n-1) \hat{x}(n-1|n-1) \right]$$

normal state transition
transforms error into a state update
expected observation
actual observation

Kalman Filter Development

Friday, April 21, 2017 8:40 AM

Given $\hat{\underline{x}}(n-1|n-1)$ and $P(n-1|n-1)$, when $y(n)$ becomes available, we want to find the unbiased estimate $\hat{\underline{x}}(n|n)$ that minimizes $\text{Tr}\{P(n|n)\}$

Proceed in two steps.

1) Given $\hat{\underline{x}}(n-1|n-1)$, find $\hat{\underline{x}}(n|n-1)$ ~ "prediction step"
- use state transition equations

2) Given $y(n)$ and $\hat{\underline{x}}(n|n-1)$, find $\hat{\underline{x}}(n|n)$

our estimate of $\underline{x}(n)$

Prediction step detail

Since $\underline{w}(n)$ is zero mean

$$\hat{\underline{x}}(n|n-1) = A(n, n-1) \hat{\underline{x}}(n-1|n-1)$$

$$\underline{e}(n|n-1) = \underline{x}(n) - \hat{\underline{x}}(n|n-1)$$

$$= A(n, n-1) \underline{x}(n-1) + \underline{w}(n) - A(n, n-1) \hat{\underline{x}}(n-1|n-1)$$

$$= A(n, n-1) \underbrace{\left[\underline{x}(n-1) - \hat{\underline{x}}(n-1|n-1) \right]}_{\underline{e}(n-1|n-1)} + \underline{w}(n)$$

if $\hat{\underline{x}}(n-1|n-1)$ is unbiased then

$$E\{\underline{e}(n|n-1)\} = A(n, n-1) E\{\underline{e}(n-1|n-1)\} + E\{\underline{w}(n)\}$$

$\hat{\underline{x}}(n-1|n-1)$ is unbiased

$\underline{w}(n)$ is zero mean

$$= 0 \quad \text{so} \quad \hat{\underline{x}}(n|n-1) \text{ is also unbiased}$$

Kalman Development cont.

Friday, April 21, 2017 9:26 AM

$$P(n|n-1) = E \left\{ \underline{e}(n|n-1) \underline{e}^T(n|n-1) \right\}$$

$$= E \left\{ \left(A(n, n-1) \underline{e}(n-1|n-1) + \underline{w}(n) \right) \left(A(n, n-1) \underline{e}(n-1|n-1) + \underline{w}(n) \right)^T \right\}$$

since $\underline{e}(n-1|n-1)$ is uncorrelated with $\underline{w}(n)$

$$P(n|n-1) = A(n, n-1) P(n-1|n-1) A^T(n, n-1) + Q_w(n)$$

Ricatti Equation

$$\hat{\underline{x}}(n|n-1) = A(n, n-1) \hat{\underline{x}}(n-1|n-1)$$

step 2
Observation

we will use a linear estimator

$$\hat{\underline{x}}(n|n) = K'(n) \hat{\underline{x}}(n|n-1) + K(n) y(n)$$

we require $\hat{\underline{x}}(n|n)$ to be unbiased $\Rightarrow E\{e(n|n)\} = 0$

and we want to choose the K' & K

to minimize $E\{\|e(n|n)\|^2\}$

$$e(n|n) = \underline{x}(n) - K'(n) \hat{\underline{x}}(n|n-1) - K(n) y(n)$$

$$= \underline{x}(n) - K'(n) \left[\underline{x}(n) - \underline{e}(n|n-1) \right] - K(n) \left[C(n) \underline{x}(n) + \underline{v}(n) \right]$$

$$= \left[I - K'(n) - K(n)C(n) \right] \underline{x}(n) + K(n) \underline{e}(n|n-1) - K(n) \underline{v}(n)$$

$$E\left\{ \downarrow \right\} = 0$$

because

$$E\{\underline{e}(n|n-1)\} = 0$$

$$E\left\{ \downarrow \right\} = 0$$

because

$$E\{\underline{v}(n)\} = 0$$

for unbiased, we must have

$$I - K'(n) - K(n)C(n) = 0$$

The Kalman Gain

Friday, April 21, 2017 9:46 AM

$K'(n) = I - K(n)C(n)$ for our estimate to be unbiased

$$\hat{x}(n|n) = [I - K(n)C(n)] \hat{x}(n|n-1) + K(n)y(n)$$

$$\hat{x}(n|n) = \hat{x}(n|n-1) + \underline{K(n)} \left[y(n) - C(n) \hat{x}(n|n-1) \right]$$

need to find $K(n)$ now

guarantees unbiasedness

$$e(n|n) = [I - K(n)C(n)] e(n|n-1) - K(n) v(n)$$

uncorrelated with $w(n)$
uncorrelated with $y(n-1)$
" with $x(n)$ and $\hat{x}(n|n-1)$

$$P(n|n) = E \{ e(n|n) e^T(n|n) \}$$

$$= [I - K(n)C(n)] P(n|n-1) [I - K(n)C(n)]^T + K(n) Q_v(n) K^T(n)$$

to minimize $\text{tr} \{ P(n|n) \}$ w.r.t. $K(n)$.

$$\frac{\partial}{\partial K} \text{tr} \{ P(n|n) \} = 0$$

$$\frac{\partial}{\partial K} \text{tr}(KA) = A^T$$

yielding ...

$$\frac{\partial}{\partial K} \text{tr}(KAK') = 2KA$$

$$[K(n)C(n) - I] P(n|n-1) C^T(n) + K(n) Q_v(n) = 0$$

solving for $K(n)$

$$K(n) = P(n|n-1) C^T(n) \left[C(n) P(n|n-1) C^T(n) + Q_v(n) \right]^{-1}$$

Updating $P(n|n)$

Friday, April 21, 2017 10:02 AM

from the previous page

$$K(n) = P(n|n-1) C^T(n) \left[C(n) P(n|n-1) C^T(n) + Q_v(n) \right]^{-1}$$

$$\left[K(n) C(n) - I \right] P(n|n-1) C^T(n) + K(n) Q_v(n) = 0$$

$$P(n|n) = \left[I - K(n) C(n) \right] P(n|n-1) \left[I - K(n) C(n) \right]^T + K(n) Q_v(n) K^T(n)$$

$$= \left[I - K(n) C(n) \right] P(n|n-1) + \underbrace{\left\{ \left[K(n) C(n) - I \right] P(n|n-1) C^T(n) + K(n) Q_v(n) \right\}}_{=0} K^T(n)$$

$$P(n|n) = \left[I - K(n) C(n) \right] P(n|n-1)$$

Summary

Wednesday, April 19, 2017 10:52 AM

state equation: $\underline{x}(n) = A(n, n-1) \underline{x}(n-1) + \underline{w}(n)$

observation equation: $y(n) = C(n) \underline{x}(n) + v(n)$

Noise statistics: $E\{\underline{w}(n) \underline{w}^T(n)\} = Q_w$

$$E\{v(n) v^T(n)\} = Q_v$$

Initialization: $\hat{\underline{x}}(0|0) = E\{\underline{x}(0)\}$

$$P(0|0) = E\{(\underline{x}(0) - E\{\underline{x}(0)\})(\underline{x}(0) - E\{\underline{x}(0)\})^T\}$$

Computation: for $n = 1, \dots$

$$\hat{\underline{x}}(n|n-1) = A(n, n-1) \hat{\underline{x}}(n-1|n-1)$$

$$P(n|n-1) = A(n, n-1) P(n-1|n-1) A^T(n, n-1) + Q_w$$

$$K(n) = P(n|n-1) C^T(n) [C(n) P(n|n-1) C^T(n) + Q_v]^{-1}$$

$$\hat{\underline{x}}(n|n) = \hat{\underline{x}}(n|n-1) + K(n) [y(n) - C(n) \hat{\underline{x}}(n|n-1)]$$

$$P(n|n) = [I - K(n) C(n)] P(n|n-1)$$

- comments:
- often we know Q_v based on characterization of our sensors
 - Q_w is often hard to measure, overestimating is better than underestimating it.
 - A large Q_w can compensate for a poor model
 - We often tune the filter by tweaking Q_w & Q_v
 - Make $P(0|0)$ very large if you are uncertain about $x(0|0)$