

The Bayes Classifier

We have been starting to look at the supervised classification problem: we are given data (\mathbf{x}_i, y_i) for $i = 1, \dots, n$, where $\mathbf{x}_i \in \mathbb{R}^d$, and $y_i \in \{1, \dots, K\}$. In this section, we suppose that we know everything there is to know about the data (in a probabilistic sense): we assume that we know the *joint distribution* of (X, Y) . If we have full knowledge of the distribution, then we can design an optimal classifier without seeing any data at all.

We now make the mathematical setup completely concrete. The “feature vector” X is a random vector¹ in \mathbb{R}^d , and the “class label” Y is a discrete random scalar in $\{1, \dots, K\}$. When we say that we have a joint probability distribution for (X, Y) , it means that we have a rule that assigns probabilities to events that obeys the Kolmogorov axioms². Given $\mathcal{X} \subset \mathbb{R}^d$ and $\mathcal{Y} \subset \{1, \dots, K\}$, the joint distribution gives us a probability

$$\mathbb{P}[X \in \mathcal{X}, Y \in \mathcal{Y}] = \text{probability that } X \text{ is in } \mathcal{X} \text{ and } Y \text{ is in } \mathcal{Y}.$$

We will treat the entries in the feature vector as continuous-valued. Fixing the feature vector X at different points \mathbf{x} results in different conditional probability mass functions (pmfs) for the class label Y :

$$\eta_k(\mathbf{x}) := p_{Y|X}(k|\mathbf{x}) = \mathbb{P}[Y = k|X = \mathbf{x}]. \quad (1)$$

The pmf $p_{Y|X}(k|\mathbf{x})$, which is also called the **a posteriori distribution**, will play a central role in much of our discussion here and throughout the course. We encounter it so often that it is useful to give it the more compact notation $\eta_k(\mathbf{x})$.

¹We use non-bold capital letters for all random variables in these notes, whether they are scalar-, vector-, matrix-, or whatever-valued.

²https://en.wikipedia.org/wiki/Probability_axioms

It is also useful to note that fixing the class label Y to different values y results in different conditional probability density functions (pdfs) for the feature vector X : $f_{X|Y}(\mathbf{x}|y)$, where

$$P[X \in \mathcal{X}|Y = y] = \int_{\mathcal{X}} f_{X|Y}(\mathbf{x}|y) d\mathbf{x}.$$

$f_{X|Y}(\mathbf{x}|y)$ is the **class conditional distribution** of X , i.e., the distribution of X given that Y belongs to class y .

A classification rule or **classifier** is simply a function $h : \mathbb{R}^d \rightarrow \{1, \dots, K\}$; that is, a function which takes a feature vector and returns a class label. We can specify this classification rule by **partitioning** \mathbb{R}^d into K regions $\Gamma_1(h), \dots, \Gamma_K(h)$, where $\Gamma_k(h)$ is the set of point that h maps to k :

$$\Gamma_k(h) = \{\mathbf{x} \in \mathbb{R}^d : h(\mathbf{x}) = k\}.$$

We will judge the quality of a classifier by the probability that it makes a mistake:

$$R(h) = P[h(X) \neq Y].$$

This is also called the **risk** of h , or the **probability of error**.

We can now ask a very well-defined question which has a clear-cut answer: What is the classifier that minimizes the probability of error? The answer is simple: given $X = \mathbf{x}$, choose the class label that maximizes the conditional probability in (1).

Theorem: Define the classifier

$$h^*(\mathbf{x}) = \arg \max_{k \in \{1, \dots, K\}} \eta_k(\mathbf{x}). \quad (2)$$

Then every other classifier h has

$$R(h) \geq R(h^*).$$

Proof: The optimality of h^* in (2) follows from carefully writing down the risk for an arbitrary classifier h , applying Bayes rule, and then showing that h^* optimizes the resulting expression. We start with an expression for $1 - R(h)$, which we will show is as *large* as possible when $h = h^*$:

$$\begin{aligned}
1 - R(h) &= \mathbb{P}[h(X) = Y] \\
&= \sum_{k=1}^K \mathbb{P}[Y = k] \cdot \mathbb{P}[h(X) = k | Y = k] \\
&= \sum_{k=1}^K \mathbb{P}[Y = k] \int_{\Gamma_k(h)} f_{X|Y}(\mathbf{x}|k) d\mathbf{x} \\
&= \sum_{k=1}^K \int_{\Gamma_k(h)} \mathbb{P}[Y = k] f_{X|Y}(\mathbf{x}|k) d\mathbf{x}.
\end{aligned}$$

By Bayes rule,

$$\eta_k(\mathbf{x}) = \frac{\mathbb{P}[Y = k] f_{X|Y}(\mathbf{x}|k)}{\sum_{\ell=1}^K \mathbb{P}[Y = \ell] f_{X|Y}(\mathbf{x}|\ell)}.$$

Note that the denominator is a function of \mathbf{x} that is independent of k ; it is in fact the marginal density $f_X(\mathbf{x})$ for X . Using this and the fact that the regions $\Gamma_k(h)$ are disjoint, we can continue the string of equalities:

$$1 - R(h) = \int_{\mathbb{R}^d} \left(\sum_{k=1}^K 1_{\Gamma_k(h)}(\mathbf{x}) f_X(\mathbf{x}) \eta_k(\mathbf{x}) \right) d\mathbf{x},$$

where $1_{\mathcal{A}}(\mathbf{x})$ is the indicator function

$$1_{\mathcal{A}}(\mathbf{x}) = \begin{cases} 1, & \mathbf{x} \in \mathcal{A}, \\ 0, & \mathbf{x} \notin \mathcal{A}. \end{cases}$$

The way we choose h^* in (2) chooses the regions so that the function inside the integral above is as large as possible; it is clear that

$$\sum_{k=1}^K 1_{\Gamma_k(h)}(\mathbf{x}) f_X(\mathbf{x}) \eta_k(\mathbf{x}) \leq \sum_{k=1}^K 1_{\Gamma_k(h^*)}(\mathbf{x}) f_X(\mathbf{x}) \eta_k(\mathbf{x}),$$

for all $\mathbf{x} \in \mathbb{R}^d$. Thus

$$\begin{aligned} 1 - R(h) &\leq \int_{\mathbb{R}^d} \left(\sum_{k=1}^K 1_{\Gamma_k(h)}(\mathbf{x}) f_X(\mathbf{x}) \eta_k(\mathbf{x}) \right) d\mathbf{x} \\ &= 1 - R(h^*), \end{aligned}$$

and so $R(h^*) \leq R(h)$.

The nearest neighbor classifier

We have just seen that the Bayes classifier is optimal. Unfortunately, it requires complete knowledge of the conditional probability mass function $\eta_k(\mathbf{x})$. In the context of machine learning, this is not a reasonable assumption. The **nearest neighbor classifier** is an *extremely* simple alternative. For any \mathbf{x} , we simply find the closest point \mathbf{x}_i in the training set and then assign \mathbf{x} the same label as its nearest neighbor.

This is an incredibly simple rule, but perhaps somewhat surprisingly we can show that as $n \rightarrow \infty$, i.e., as the size of our training data grows, this simple classifier is near-optimal. To see this, we will consider the risk of the nearest neighbor classifier h^{NN} conditioned on $X = \mathbf{x}$ and compare this to the risk of the Bayes classifier h^* .

To make our discussion simpler, we will restrict our attention to the case of binary classification where $y_i \in \{0, 1\}$. We first note that the

risk of the Bayes classifier h^* conditioned on $X = \mathbf{x}$ is given by

$$R^*(\mathbf{x}) := \mathbb{P}[Y \neq h^*(\mathbf{x}) | X = \mathbf{x}].$$

If $h^*(\mathbf{x}) = 0$ then we have $R^*(\mathbf{x}) = \mathbb{P}[Y = 1 | X = \mathbf{x}] = \eta_1(\mathbf{x})$. Similarly, if $h^*(\mathbf{x}) = 1$ we have $R^*(\mathbf{x}) = \eta_0(\mathbf{x})$. Since by definition $h^*(\mathbf{x})$ selects the label that *maximizes* $\eta_k(\mathbf{x})$, we thus have that

$$R^*(\mathbf{x}) = \min\{\eta_0(\mathbf{x}), \eta_1(\mathbf{x})\}. \quad (3)$$

For the nearest neighbor classifier, note that

$$R^{\text{NN}}(\mathbf{x}) := \mathbb{P}[h^{\text{NN}}(\mathbf{x}) \neq Y | X = \mathbf{x}].$$

In our analysis, we will treat not only (X, Y) as random, but also the output $h^{\text{NN}}(\mathbf{x})$ as random since it depends on the dataset, which is itself drawn at random from the same distribution as (X, Y) . This allows us to write

$$\begin{aligned} R^{\text{NN}}(\mathbf{x}) &= \mathbb{P}[Y = 0 | X = \mathbf{x}] \mathbb{P}[h^{\text{NN}}(\mathbf{x}) = 1 | X = \mathbf{x}] \\ &\quad + \mathbb{P}[Y = 1 | X = \mathbf{x}] \mathbb{P}[h^{\text{NN}}(\mathbf{x}) = 0 | X = \mathbf{x}]. \end{aligned} \quad (4)$$

If \mathbf{x}_{NN} denotes the nearest neighbor to \mathbf{x} , then we can write

$$\mathbb{P}[h^{\text{NN}}(\mathbf{x}) = k | X = \mathbf{x}] = \mathbb{P}[Y = k | X = \mathbf{x}_{\text{NN}}] = \eta_k(\mathbf{x}_{\text{NN}}).$$

As $n \rightarrow \infty$, we have that $\|\mathbf{x}_{\text{NN}} - \mathbf{x}\| \rightarrow 0$, and thus as $n \rightarrow \infty$ we have

$$\eta_k(\mathbf{x}_{\text{NN}}) \rightarrow \eta_k(\mathbf{x}).$$

Plugging this back into (4) and simplifying, we obtain

$$\begin{aligned} R^{\text{NN}}(\mathbf{x}) &\rightarrow \eta_0(\mathbf{x})\eta_1(\mathbf{x}) + \eta_1(\mathbf{x})\eta_0(\mathbf{x}) \\ &= 2\eta_0(\mathbf{x})\eta_1(\mathbf{x}) \\ &\leq 2 \min\{\eta_0(\mathbf{x}), \eta_1(\mathbf{x})\}, \end{aligned}$$

where the last inequality follows from the fact that both $\eta_1(\mathbf{x})$ and $\eta_2(\mathbf{x})$ are less than 1. Combining this with (3), this yields

$$\lim_{n \rightarrow \infty} R^{\text{NN}}(\mathbf{x}) \leq 2R^*(\mathbf{x}),$$

or in words, that asymptotically, the risk of the nearest neighbor classifier is at most twice the Bayes risk.

This can be strengthened by considering the more general k -nearest neighbors classifier. The idea here is to assign a label to \mathbf{x} by taking a majority vote over the k training points closest to \mathbf{x} . If $R^{\text{kNN}}(\mathbf{x})$ denotes the risk of the k -nearest neighbor classifier, then one can show via a similar argument that

$$\lim_{n \rightarrow \infty} R^{\text{kNN}}(\mathbf{x}) \leq \left(1 + \sqrt{2/k}\right) R^*(\mathbf{x}).$$

Thus, by increasing k it is possible to drive this multiplicative constant arbitrarily close to 1. This results in a property known as **universal consistency**. Specifically, if R^* denotes the Bayes risk and R_n^{kNN} denotes the risk of the k -nearest neighbors classifier based on a dataset of size n , then one can show that as $n \rightarrow \infty$, if $k \rightarrow \infty$ while $k/n \rightarrow 0$, then $R_n^{\text{kNN}} \rightarrow R^*$.

In words this is simply saying that for any possible distribution on the data, if we are given enough data eventually the risk of the k -nearest neighbor classifier will converge to the Bayes risk (i.e., to the optimal risk). Unfortunately (or fortunately, depending on your perspective), you might have to wait a very long time, so there is still a role for other machine learning algorithms to improve on this situation when we only have a finite amount of data.